

ECE276A: Sensing & Estimation in Robotics

Lecture 2: Probability Theory

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Project 1 is out

- ▶ Click on [P1](#) on the schedule and enter the password
- ▶ **Start early!** We have not covered everything but:
 - ▶ Make sure you can open and download everything
 - ▶ Try loading, viewing, and labeling the data
 - ▶ Work on the first 2 problems
 - ▶ Practice your python skills
 - ▶ Start your report
- ▶ TA sessions: Friday, 2:00 - 3:30 pm, Jacobs Hall 2315
- ▶ First TA session – python tutorial

Events

- ▶ **Experiment:** any procedure that can be repeated infinitely and has a well-defined set of possible outcomes.
- ▶ **Sample space Ω :** the set of possible outcomes of an experiment.
 - ▶ $\Omega = \{HH, HT, TH, TT\}$
 - ▶ $\Omega = \{\square, \square, \square, \square, \square, \square\}$
- ▶ **Event A :** a subset of the possible outcomes Ω
 - ▶ $A = \{HH\}$, $B = \{HT, TH\}$
- ▶ **Probability of an event:** $\mathbb{P}(A) = \frac{N_A}{N} = \frac{\text{\#possible occurrences of } A}{\text{\#all possible outcomes}}$

Probability Axioms

▶ Probability Axioms:

- ▶ $\mathbb{P}(A) \geq 0$
- ▶ $\mathbb{P}(\Omega) = 1$
- ▶ If $\{A_i\}$ are disjoint ($A_i \cap A_j = \emptyset$), then $\mathbb{P}(\bigcup_i A_i) = \sum_i \mathbb{P}(A_i)$

▶ Corollary

- ▶ $\mathbb{P}(\emptyset) = 0$
- ▶ $\max\{\mathbb{P}(A), \mathbb{P}(B)\} \leq \mathbb{P}(A \cup B) = \mathbb{P}(A) + \mathbb{P}(B) - \mathbb{P}(A \cap B) \leq \mathbb{P}(A) + \mathbb{P}(B)$
- ▶ $A \subseteq B \Rightarrow \mathbb{P}(A) \leq \mathbb{P}(B)$

Set of Events

- ▶ **Conditional Probability:** $\mathbb{P}(A | B) = \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)}$, $\mathbb{P}(B) \neq 0$
- ▶ **Total Probability Theorem:** If $\{A_1, \dots, A_n\}$ is a partition of Ω , i.e., $\Omega = \bigcup_i A_i$ and $A_i \cap A_j = \emptyset, i \neq j$, then:

$$\mathbb{P}(B) = \sum_{i=1}^n \mathbb{P}(B \cap A_i)$$

- ▶ **Bayes Theorem** If $\{A_1, \dots, A_n\}$ is a partition of Ω , then:

$$\mathbb{P}(A_i | B) = \frac{\mathbb{P}(B | A_i)\mathbb{P}(A_i)}{\sum_{j=1}^n \mathbb{P}(B | A_j)\mathbb{P}(A_j)}$$

- ▶ **Independent events:** $\mathbb{P}(\bigcap_i A_i) = \prod_i \mathbb{P}(A_i)$
 - ▶ observing one does not give any information about another
 - ▶ in contrast, disjoint events never occur together: one occurring tells you that others will not occur and hence, disjoint events are always dependent

Measure Space

- ▶ **σ -algebra**: a collection of subsets of Ω closed under complementation and countable unions.
- ▶ **Measurable space**: a tuple (Ω, \mathcal{F}) , where Ω is a sample space and \mathcal{F} is a σ -algebra.
- ▶ **Measure**: a function $\mu : \mathcal{F} \rightarrow \mathbb{R}$ satisfying $\mu(A) \geq \mu(\emptyset) = 0$ for all $A \in \mathcal{F}$ and countable additivity $\mu(\cup_i A_i) = \sum_i \mu(A_i)$ for disjoint A_i .
- ▶ A measure μ is **σ -finite** on (Ω, \mathcal{F}) , if Ω can be obtained as the countable union $\cup_n A_n$ of sets $A_n \in \mathcal{F}$ of finite measure, $\mu(A_n) < \infty$.

Probability Space

- ▶ **Probability measure:** a measure that satisfies $\mu(\Omega) = 1$.
- ▶ **Probability space:** a triple $(\Omega, \mathcal{F}, \mathbb{P})$, where Ω is a sample space, \mathcal{F} is a σ -algebra, and $\mathbb{P} : \mathcal{F} \rightarrow [0, 1]$ is a probability measure.
- ▶ **Borel σ -algebra \mathcal{B} :** the smallest σ -algebra containing all open sets from a topological space. Necessary because there is no valid translation invariant way to assign a finite measure to all subsets of $[0, 1]$.

Random Variable

- ▶ **Random variable** X : an \mathcal{F} -measurable function from (Ω, \mathcal{F}) to $(\mathbb{R}, \mathcal{B})$, i.e., a function $X : \Omega \rightarrow \mathbb{R}$ s.t. the preimage of every set in \mathcal{B} is in \mathcal{F} .
- ▶ **Distribution function** $F(x)$ of a random variable X : a function $F(x) := \mathbb{P}(X \leq x)$ that is non-decreasing, right-continuous, and $\lim_{x \rightarrow \infty} F(x) = 1$ and $\lim_{x \rightarrow -\infty} F(x) = 0$.

- ▶ **Density/mass function** $f(x)$ of a random variable X

Continuous RV

Discrete RV

$X : (\Omega, \mathcal{F}, \mathbb{P}) \rightarrow (\mathbb{R}, \mathcal{B}, \mathbb{P} \circ X^{-1})$:

$X : (\Omega, 2^\Omega, \mathbb{P}) \rightarrow (\mathbb{R}, \mathcal{B}, \mathbb{P} \circ X^{-1})$:

- ▶ $f(x) \geq 0$

- ▶ $f(x) = \mathbb{P}(X = x) \geq 0$

- ▶ $\int f(y) dy = 1$

- ▶ $\sum_i f(i) = 1$

- ▶ $F(x) = \int_{-\infty}^x f(y) dy = \mathbb{P}(X \leq x)$

- ▶ $F(x) = \sum_{i \in \mathbb{Z}, i \leq x} f(i) = \mathbb{P}(X \leq x)$

- ▶ $\mathbb{P}(X = x) = F(x) - F(x^-) = \lim_{\epsilon \rightarrow 0} \int_{x-\epsilon}^x f(y) dy = 0$

- ▶ $\mathbb{P}(a < X \leq b) = F(b) - F(a) = \int_a^b f(x) dx$

Expectation

- ▶ **Lebesgue Integration:** The integral $\int g d\mu$ of a measurable function g on measurable space (Ω, \mathcal{F}) with a σ -finite measure μ can be defined. In the case that μ has a pdf p , the Lebesgue integral is equivalent to a Riemann integral: $\int g d\mu = \int g(x)p(x)dx$.
- ▶ **Expectation:** Given a random variable $X: (\Omega, \mathcal{F}, \mathbb{P}) \rightarrow (\mathbb{R}^n, \mathcal{B}^n, \mathbb{P} \circ X^{-1})$ and a measurable function $g: (\mathbb{R}^n, \mathcal{B}^n, \mathbb{P} \circ X^{-1}) \rightarrow (\mathbb{R}^m, \mathcal{B}^m, \mathcal{L})$, the expectation of $g(X)$ is defined as follows:

$$\mathbb{E}[g(X)] = \int_{\Omega} g(X(\omega)) d\mathbb{P}(\omega) = \int_{\mathbb{R}^n} g(x) d\mathbb{P}(X^{-1}(x)) = \int_{\mathbb{R}^m} y d\mathcal{L}(y)$$

When X has a pdf p and g has a pdf l , the above simplifies to:

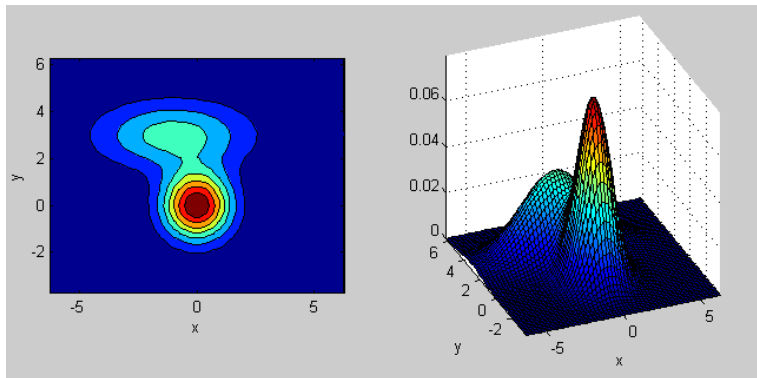
$$\mathbb{E}[g(X)] = \int_{\mathbb{R}^n} g(x)p(x)dx = \int_{\mathbb{R}^m} yl(y)dy$$

- ▶ **Expectation of an Indicator:** $\mathbb{E}[\mathbb{1}_A] = \int \mathbb{1}_A(\omega) d\mathbb{P}(\omega) = \mathbb{P}(A)$
- ▶ **Variance of a random variable X :**
 $Var[X] := \mathbb{E}[(X - \mathbb{E}[X])(X - \mathbb{E}[X])^T] = \mathbb{E}[XX^T] - \mathbb{E}[X]\mathbb{E}[X]^T$

Gaussian Distribution

- ▶ The **Mahalaonobis distance** for vector $x \in \mathbb{R}^n$ and symmetric positive-definie matrix $S \in \mathbb{S}_{>0}^n$ is: $\|x\|_S^2 := x^T S^{-1} x$
- ▶ **Gaussian random variable** $X \sim \mathcal{N}(\mu, \Sigma)$
 - ▶ paramteres: **mean** $\mu \in \mathbb{R}^n$, **covariance** $\Sigma \in \mathbb{S}_{\geq 0}^n$
 - ▶ pdf: $\phi(x; \mu, \Sigma) := \frac{1}{\sqrt{(2\pi)^n \det(\Sigma)}} \exp(-\frac{1}{2}(x - \mu)^T \Sigma^{-1}(x - \mu))$
 - ▶ expectation: $\mathbb{E}[X] = \int x \phi(x; \mu, \Sigma) dx = \mu$
 - ▶ variance: $\text{Var}[X] = \Sigma$
- ▶ **Gaussian mixture** $X \sim \mathcal{NM}(\{\alpha_k\}, \{\mu_k\}, \{\Sigma_k\})$
 - ▶ parameters: **weights** $\alpha_k \geq 0$, $\sum_k \alpha_k = 1$,
means $\mu_k \in \mathbb{R}^n$, **covariances** $\Sigma_k \in \mathbb{S}_{\geq 0}^n$
 - ▶ pdf: $p(x) := \sum_k \alpha_k \phi(x; \mu_k, \Sigma_k)$
 - ▶ expectation: $\mathbb{E}[X] = \int x p(x) dx = \sum_k \alpha_k \mu_k =: \bar{\mu}$
 - ▶ variance: $\mathbb{E}[X X^T] - \mathbb{E}[X] \mathbb{E}[X]^T = \sum_k \alpha_k (\Sigma_k + \mu_k \mu_k^T) - \bar{\mu} \bar{\mu}^T$

Mixture of two 2-D Gaussians



Other Distributions

- ▶ **Uniform continuous random variable** $X \sim \mathcal{U}(a, b)$
 - ▶ parameters: $-\infty < a < b < \infty$
 - ▶ pdf: $p(x) = \frac{1}{b-a}$ for $x \in [a, b]$
 - ▶ expectation: $\mathbb{E}[X] = \frac{1}{2}(a + b)$
 - ▶ variance: $\text{Var}[X] = \frac{1}{12}(b - a)^2$

- ▶ **Chi-Square random variable** $X \sim \chi^2(k)$
 - ▶ parameters: degrees of freedom $k \in \mathbb{N}$
 - ▶ pdf: $p(x) = \frac{1}{2^{k/2}\Gamma(\frac{k}{2})}x^{\frac{k}{2}-1}e^{-\frac{x}{2}}$ for $x \geq 0$
 - ▶ expectation: $\mathbb{E}[X] = k$
 - ▶ variance: $\text{Var}[X] = 2k$
 - ▶ if $Y \sim \mathcal{N}(\mu, \Sigma)$, then $X := \|Y - \mu\|_{\Sigma}^2 \sim \chi^2(n)$

- ▶ The **Gamma function** $\Gamma(\alpha) := \int_0^{\infty} y^{\alpha-1}e^{-y}dy$, $\alpha > 0$, satisfies $\Gamma(\frac{1}{2}) = \sqrt{\pi}$, $\Gamma(1) = 1$, $\Gamma(k + 1) = k\Gamma(k)$

Composition and Norm

- ▶ **Change of Density:** Let $Y = f(X)$. Then with $dy = \det\left(\frac{df}{dx}(x)\right) dx$:

$$\mathbb{P}(Y \in A) = \mathbb{P}(X \in f^{-1}(A)) = \int_{f^{-1}(A)} p_x(x) dx$$

$$\underbrace{\text{change of}}_{\text{variables}} \int_A \boxed{\frac{1}{\det\left(\frac{df}{dx}(f^{-1}(y))\right)} p_x(f^{-1}(y))} dy$$

- ▶ **L^p -Space:** Let (S, Σ, μ) be a measure space and $1 \leq p < \infty$. Then, $f \in L^p(\mu)$ if f is measurable wrt S and

$$\boxed{\|f\|_p := \left(\int_S |f|^p d\mu \right)^{\frac{1}{p}} < \infty}$$

- ▶ **Differential Entropy** of a continuous random variable X with pdf p is:

$$h(p) := - \int p(x) \log p(x) dx$$

- ▶ **Kullback-Leibler (KL) divergence** from pdf p to pdf q is:

$$d_{\mathcal{KL}}(p||q) := \int p(x) \log \frac{p(x)}{q(x)} dx$$

Example: Change of Density

- ▶ Let $V := (X, Y)$ be a random vector with pdf:

$$p_V(x, y) := \begin{cases} 2y - x & x < y < 2x \text{ and } 1 < x < 2 \\ 0 & \text{else} \end{cases}$$

- ▶ Let $T := (M, N) = g(V) := \left(\frac{2X-Y}{3}, \frac{X+Y}{3}\right)$ be a function of V
- ▶ Note that $X = M + N$ and $Y = 2N - M$ and hence the pdf of V is non-zero for $0 < m < n/2$ and $1 < m + n < 2$. Also:

$$\det\left(\frac{dg}{dv}\right) = \det\begin{bmatrix} 2/3 & -1/3 \\ 1/3 & 1/3 \end{bmatrix} = \frac{1}{3}$$

- ▶ The pdf T is:

$$p_T(m, n) = \begin{cases} \frac{1}{\det\left(\frac{dg}{dv}(m+n, 2n-m)\right)} p_V(m+n, 2n-m), & 0 < m < n/2 \text{ and} \\ & 1 < m+n < 2, \\ 0, & \text{else.} \end{cases}$$

Inequalities

▶ **Markov's Inequality**

$$0 \leq X, f$$

\Rightarrow

$$\mathbb{P}(X \in A) \leq \frac{\mathbb{E}f(X)}{\inf_{x \in A} f(x)}$$

Examples:

$$\mathbb{P}(X \geq a) \leq \frac{\mathbb{E}X}{a}$$

and

$$\mathbb{P}(|X - \mathbb{E}X| \geq a) \leq \frac{\text{Var}(X)}{a}$$

▶ $f : \mathbb{R}^n \rightarrow \mathbb{R}$ is **convex** if one of the following holds:

- ▶ $f(\lambda x + (1 - \lambda)y) \leq \lambda f(x) + (1 - \lambda)f(y), \quad \forall x, y \in \mathbb{R}^n, \lambda \in [0, 1]$
- ▶ $f(y) \geq f(x) + \nabla f(x)^T (y - x), \quad \forall x, y \in \mathbb{R}^n$
- ▶ $\nabla^2 f(x) \succeq 0, \quad \forall x \in \mathbb{R}^n$

▶ **Jensen's Inequality** Let X be random variable and $\psi : \mathbb{R} \rightarrow \mathbb{R}$ be a convex function. Then, $\mathbb{E}\psi(X) \geq \psi(\mathbb{E}(X))$ provided both exist.

- ▶ $|\mathbb{E}|X_t| - \mathbb{E}|X|| \leq \mathbb{E}||X_t| - |X|| \leq \mathbb{E}|X_t - X|$, thus convergence in L^1 implies convergence in expectation but *not the converse*
- ▶ $(\mathbb{E}|X|)^2 \leq \mathbb{E}X^2$
- ▶ $f\left(\frac{\sum_i a_i z_i}{\sum_i a_i}\right) \leq \frac{\sum_i a_i f(z_i)}{\sum_i a_i}$ for convex f , points $\{z_i\}$ in f 's domain, and $a_i \geq 0$

Inequalities

- ▶ **Hölder's Inequality** (generalizes **Cauchy-Schwarz**) Let (S, Σ, μ) be a measure space and let $p, q \in [1, \infty]$ with $1/p + 1/q = 1$. Then, for

measurable f, g on S : $\|fg\|_1 \leq \|f\|_p \|g\|_q$.

- ▶ If $f, g \in L^2(\mu)$, then $|\langle f, g \rangle| \leq \|f\|_2 \|g\|_2$
- ▶ $\mathbb{E}|XY| \leq (\mathbb{E}|X|^p)^{\frac{1}{p}} (\mathbb{E}|Y|^q)^{\frac{1}{q}}$
- ▶ $\mathbb{E}(|X|^r) \leq (\mathbb{E}(|X|^s))^{\frac{r}{s}}$
- ▶ $\sum_{k=1}^{\infty} |x_k y_k| \leq \left(\sum_{k=1}^{\infty} |x_k|^p \right)^{\frac{1}{p}} \left(\sum_{k=1}^{\infty} |y_k|^q \right)^{\frac{1}{q}}$
- ▶ $|a^T b| \leq \|a\|_2 \|b\|_2$

Set of Random Variables

- ▶ The **joint distribution** of random variables $\{X_i\}_{i=1}^n$ on $(\Omega, \mathcal{F}, \mathbb{P})$ defines their simultaneous behavior and is associated with a cumulative distribution function $F(x_1, \dots, x_n) := \mathbb{P}(X_1 \leq x_1, \dots, X_n \leq x_n)$. The CDF $F_i(x_i)$ of X_i defines its **marginal distribution**.
- ▶ Random variables $\{X_i\}_{i=1}^n$ on $(\Omega, \mathcal{F}, \mathbb{P})$ are **jointly independent** iff for all $\{A_i\}_{i=1}^n \subset \mathcal{F}$, $\mathbb{P}(X_i \in A_i, \forall i) = \prod_{i=1}^n \mathbb{P}(X_i \in A_i)$
- ▶ Let X and Y be random variables and suppose $\mathbb{E}X$, $\mathbb{E}Y$, and $\mathbb{E}XY$ exist. Then, X and Y are **uncorrelated** iff $\mathbb{E}XY = \mathbb{E}X\mathbb{E}Y$ or equivalently $\text{Cov}(X, Y) = 0$.
- ▶ Independence implies uncorrelatedness
- ▶ Two random variables X and Y are orthogonal if $\mathbb{E}[X^T Y] = 0$

Set of Random Variables

- ▶ **Total Probability Theorem** Given two random variables X_1, X_2 with a joint pdf p , one can obtain the marginal pdf p_1 of X_1 as follows:
$$p_1(x_1) := \int p(x_1, x_2) dx_2.$$
- ▶ **Conditional expectation** Let X be an RV on $(\Omega, \mathcal{F}_0, \mathbb{P})$ with $\mathbb{E}|X| < \infty$ and let $\mathcal{F} \subseteq \mathcal{F}_0$. Then, $Y := \mathbb{E}[X | \mathcal{F}]$ is an RV that satisfies:
 - ▶ (Measurability Axiom) $Y \in \mathcal{F}$,
 - ▶ (Integral Axiom) $\int_G Y d\mathbb{P} = \int_G X d\mathbb{P}$ for all $G \in \mathcal{F}$.

Y exists and is unique up to values on a set of measure zero. The following notation is common:

 - ▶ $\mathbb{E}[X | Z] := \mathbb{E}[X | \sigma(Z)]$
 - ▶ $\mathbb{P}(A | B) := \mathbb{E}[\mathbb{1}_A | \sigma(B)]$

Set of Random Variables

- ▶ **Conditional distribution** If (X, Y) has a pdf f on \mathbb{R}^2 and $\mathbb{E}|g(X)| < \infty$, then $\mathbb{E}[g(X) | \sigma(Y)] = h(Y)$ for $h(y) := \int g(x) \frac{f(x,y)}{\int f(x,y) dx} dx$. Note that this defines the pdf of X

conditioned on $Y = y$ as $p(x|y) := \frac{f(x,y)}{\int f(x,y) dx}$

- ▶ **Bayes Theorem** The conditional, marginal, and joint pdfs of X and Y are related:

$$p(x, y) = p(y|x)p(x) = p(x|y)p(y)$$

$$\Rightarrow p(x|y) = \frac{p(y|x)p(x)}{\int p(y | x')p(x') dx'}$$

Set of Random Variables

- ▶ **Convolution** Let X and Y be independent random variables with pdfs f and g , respectively. Then, the pdf of $Z = X + Y$ is given by the convolution of f and g :

$$[f * g](z) := \int f(z - y)g(y)dy$$

- ▶ **Variance**
$$\text{Var} \left(\sum_{i=1}^n X_i \right) = \sum_{i=1}^n \text{Var}(X_i) + \sum_{i=1}^n \sum_{j \neq i} \text{Cov}(X_i, X_j)$$

$$\text{Cov}(X_i, X_j) := \mathbb{E} \left((X_i - \mathbb{E}X_i)(X_j - \mathbb{E}X_j)^T \right) = \mathbb{E}(X_i X_j^T) - \mathbb{E}X_i \mathbb{E}X_j^T$$